

# SHOULD FIRMS STRIVE FOR THE EDUCATIONAL DIVERSITY OF THE WORKFORCE? ESTIMATION OF THE IMPACT OF FIRMS' EDUCATIONAL STRUCTURE ON SALES GROWTH AND EXPORTS

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## Abstract

This study analyses the relationship between a firms' growth and the educational diversity of their workforce. It differentiates itself from other studies by using two measures of human capital diversity – one for education levels and one for fields – instead of one. The results show that workers with more diverse specializations than the workforce of competitors benefit the firm's growth and extensive margin of trade. Contrastingly, above-average workforce heterogeneity in education levels does not seem to affect firms' performance. Furthermore, it is shown that the association between human capital diversity and firms' performance differs among sectors. Compared to previous studies, which concentrated on one particular industry or one measure of educational diversity, this more comprehensive approach allows reconciling some of the contradicting results. The study shows that the mixed results in existing papers can result from dissimilar impacts of various measures of human capital diversity on firm performance or its differences among sectors.

Keywords: education levels; education fields; diversity; firm growth; human capital; exports; workforce structure



## INTRODUCTION

The importance of human capital for firm performance has become a fact in economic literature. Already in the mid-twenty century, Becker (1964) defined human capital as skills and knowledge that individuals obtain through investments in schooling, on-the-job training, and other types of experience. As such, it can be considered as an additional input (separate from labour) to the production function (e.g. Black & Lynch, 1996) and one of the factors that determine firm growth (for a review of all determinants, see Coad, 2009; Davidsson, Achtenhagen, & Naldi, 2010).

Two competing theories in a special brand of literature explain the relationship between human capital and firm performance in international trade. According to the first, there are fixed costs of exporting. Only firms with higher endowments of resources (including human capital) and hence productivity can 'invest' in exporting (Clerides et al., 1998; Bernard & Jensen, 1999; Melitz, 2003). However, this productivity effect is not likely to persist among exporters, so the more productive firms are not likely to export more than less productive firms after becoming exporters. The competing theory with less empirical support claims that firms become more productive (and increase human capital) due to the competition and learning in international markets (e.g. Clerides et al., 1998).

Furthermore, theory shows that the level of human capital and its diversity affect firm growth through productivity, and the relationship between the two can be either positive or negative. Using a theoretical model that follows the production function approach, Lazear (1999) argued that the gains from workforce diversity are most significant when individuals can communicate and understand each other but have entirely disjoint skills and information sets, which are all relevant for the tasks that have to be performed within the firm. The relation between human capital diversity and firm productivity will be positive if there is sufficient mutual learning and collaboration among workers with different human capital (Hamilton, Nickerson, & Owan, 2012). According to theory, the impact of human capital heterogeneity also depends on several firm characteristics—it is more beneficial to firms that depend on the innovations (Jehn, Northcraft, & Neale, 1999; Prat, 2002) and have greater complexity of tasks (Jehn, 1995; Jehn et al., 1999; Parrotta, Pozzoli, & Pytlikova, 2014b; Stewart, 2006). Nevertheless, diversity can also hamper job satisfaction, communication and possibilities of a confrontation between competing views, which could negatively affect firm performance (Festinger, 1954; Grund & Westergaard-Nielsen, 2008; Pfeffer, 1985).

Empirical studies have not yet reached a consensus on the relationship between the educational diversity of the workforce and firm performance. Although firms' educational heterogeneity of human capital can stem either from differences in levels or fields of education, the existing studies concentrated only on one of them or the combination of both. But, do the increases of the two measures of educational workforce diversity have equivalent impacts on firms' performance? If not, the different educational heterogeneity measures might explain at least a part of the contrasting empirical results. The other part of the discrepancies in the empirical results could arise due to the different firm performance measures or differences in the industry of the firms selected in the samples. Here presented analysis tackles the question above by analyzing the impact of both measures of educational heterogeneity on growth, the probability to export and export intensity across diverse industries.

Authors have thus far measured firm performance with value added (Ganero et al., 2014; Navon, 2010; Irnazo et al., 2008; Ilmakunnas & Ilmakunnas, 2011), total factor productivity (Parrotta et al., 2014a; Ilmakunnas & Ilmakunnas, 2011), turnover rate (Jackson et al., 1991; Tsui et al., 1992; Hirsch et al., 2020), growth (Smith et al., 1994; Hambrick et al., 1996; Simons et al., 1999), worker/team performance (Jehn & Bezrukova, 2004; Kearney & Gebert, 2009; Kurtulus, 2011), firm survival (Backman & Kohlhase, 2020), profit (Lee & Kim, 2019), firm's innovation (Parrotta et al., 2014b), efficiency (Vanderberghe, 2016) and strategic consensus (Knight et al., 1999). The majority of these studies



measured education diversity only in terms of levels of education. In contrast, Parotta et al. (2014a, 2014b), Irnazo et al. (2008), Jackson et al. (1991) and Backman and Kohlhase (2020) used one combined measure of diversity in education levels and fields, while Navon (2010), Hambrick et al. (1996) and Kearney and Gebert (2009) concentrated only on workforce diversity in fields of education. None of these studies differentiated between diversity in education levels and diversity in education fields. Out of the 19 above-listed studies, only 7 used a sample of firms from more than one sector; 12 of them found a positive effect on performance, 4 found a negative effect on performance, and 3 an insignificant one.

The empirical evidence on firm performance in international trade provided more support for the theory with fixed costs of entering export markets (e.g., Munch & Skaksen, 2008; Ganotakis & Love, 2011; Eliasson et al., 2012; Mulliqi et al. 2019). Papers measuring human capital with shares of high skilled or low skilled workers provided evidence for a positive link between exports and the workforce's education (Ganotakis & Love, 2011; Wagner, 2012; Brambilla, 2017; Mulliqi et al., 2019; Chiappini, 2021). However, the effect might differ between sectors (Mulliqi et al., 2019), types of products, and countries' income level (Luong & Chen, 2016). The impact of educational diversity of the workforce on exports has been underexplored thus far.

This paper fills the literature gap by investigating how the diversity of all employees with respect not only to education levels but also fields of education affects the growth of total sales and foreign sales per employee and the probability of being an exporter in a wide variety of industries.<sup>1</sup> In addition, it examines whether estimates of diversity–growth nexus differ between knowledge-intensive or high-tech firms and those that are not. This more comprehensive approach, which is not concentrated on one particular industry or one measure, is possible due to the unique register-based linked employer-employee dataset from Slovenia and allows to reconcile some of the previous contradicting results. As such, the analysis is undoubtedly relevant for the international audience.

Two separate measures of human capital diversity – one for levels of education and one for fields – reveal that increase in diversity of fields of education above the one of competing firms is positively associated with firm growth. However, the same does not hold for diversity in levels of education. The diversity in education fields also impacts the extensive margin of trade, but no significant effects were found for the intensive margin. The results, therefore, suggest that firms should strive for human capital diversity. However, the workforce should be heterogeneous in terms of fields rather than levels of education.

The paper continues with a short description of the institutional framework, which is followed by a portrayal of empirical strategy and characterization of the dataset. The following section presents results and discussion. The last section concludes the paper.

## **A VERY SHORT DESCRIPTION OF THE SLOVENIAN ECONOMY**

Slovenia is a small, open economy whose GDP per capita was 20,810 EUR in 2017 (69.2 % of EU28 total GDP per capita). Its real GDP growth rate in 2017 equalled 4.8, exceeding the EU28 average by 2.2 percentage points. 20.6% of the total value added was created in manufacturing. The second-largest sector by created value added was wholesale and retail trade, which generated 10.4% of the total value added. Exports of goods and services in the year 2017 presented as much as 83.2 % of GDP — highly exceeding the EU28 average (45.7%) —, while imports presented 74.3% of GDP (EU28 average was 42.2 %). The main Slovenian trade partners were Germany, Italy, Austria, and Croatia. Exports to these countries presented 47% of all exports and 51% of all imports (Eurostat, 2020a; SURS, 2020).

The majority of 195,756 Slovenian firms were micro or small enterprises – there were only 2,084 medium-sized firms and 346 large firms with more than 250 employees in 2017. The majority of large firms operated in the manufacturing sector. Besides manufacturing, the most important



economic activities by the number of firms were professional, scientific, and technical activities (NACE section M), wholesale and retail trade (NACE section G), construction (NACE section F) and other service activities (NACE section S) (SURS, 2020). Slovenia had roughly 845 thousand active persons in 2017 and a 6.6% unemployment rate. 56.2 % of active persons had secondary education, and 31.4% per cent were tertiary educated. Females presented 41.7% of the active labour force.

## EMPIRICAL STRATEGY

Traditionally, the growth theory of the firms is empirically verified with a dynamic panel regression model. The specification, which takes into account the firm's micro and macro factors, can be written as:

$$y_{it} = \alpha y_{it-1} + \beta X_{it-1} + \gamma Z_{it-1} + \mu_i + \tau_t + \varepsilon_{it}, \quad (1)$$

$y$  presents one of the variables of primary interest: growth in sales per employee, growth in foreign sales per employee, or probability of exporting;  $X$  is a vector of microeconomic factors that potentially impact performance (e.g. size, age, financial constraints, human capital);  $Z$  stands for macroeconomic factors such as industry, business environment or tightness of the labour market;  $\mu$  and  $\tau$  represent firm-specific fixed effects and time effects, respectively.

The main variables of interest are human capital diversity indicators with respect to the level and field of education. The diversity indicators are calculated with a normalized Herfindahl index (NHI):

$$d_{i,k} = 1 - NHI_{i,k} = 1 - \frac{\sum_{i=1}^N s_{i,k}^2 - \frac{1}{N}}{1 - \frac{1}{N}}, \quad (2)$$

where  $s_{i,k}$  is a share of employees with specific characteristics (e.g. specific level of education) in a firm  $i$  of a specific size in industry  $k$ , and  $N$  is the number of groups defined by these characteristics (e.g. the number of levels of education). Diversity in each dimension thus ranges between 0 and 1, with lower values indicating lower diversity and vice versa. However, the levels of diversity that are the most beneficial to the firm's performance might depend on firm size and industry. For example, it might be optimal for a small programming firm to employ only highly skilled programmers. However, a middle-sized programming firm might struggle without employees with managerial and marketing skills. In order to eliminate such size and industry-specific effects, the equations include normalized diversity indicators:

$$DI_i = d_{i,k} - \bar{d}_k, \quad (3)$$

where  $\bar{d}_k$  presents an average  $d_i$  of firms of particular size and industry ( $k$ ). The regression coefficients

for normalized diversity indicators thus measure the association between the firm's performance and deviation from average diversity in firms of the same size and industry. A positive estimated parameter will therefore indicate that an increase in human capital diversity relative to the average diversity in similar firms is positively related to the dependent variable, i.e. growth in sales per employee or



probability of exporting. Similarly, the level of human capital is measured with the normalized average education, which is calculated as the difference between the average education level of a firm and the average education level of similar firms in terms of size and industry.

Since unobserved firm-specific time-invariant characteristics such as ownership of a patent, access to market, or resources can cause omitted variable bias, they must be removed by an appropriate econometric estimation technique. Due to inconsistency of within and first-difference estimator in autoregressive panel-data models, the estimation is performed with panel GMM estimation (Arellano & Bond, 1991; Arellano & Bover, 1995; Blundell & Bond, 1998) and followed by two specification tests: Sargan test of overidentifying restrictions and Arellano-Bond test for zero autocorrelation in the first differenced errors.

## **DATA AND DESCRIPTIVE STATISTICS**

The empirical analysis uses matched employer-employee dataset of private firms operating in Slovenia in economic activities coded B–N (without L) by NACE classification in the 2008–2017 period. As the aim is to analyze the impact of human capital diversity, the sample includes firms in the form of general partnerships, limited partnerships, limited liability partnerships, or public limited companies with more than 10 employees that have operated at least three years during 2008–2017 (but not necessarily the entire observed period).

Table 1 presents the means and standard deviations of selected variables. Variable education level can take four values: 1 for employees with primary education or lower, 2 for those with secondary education, 3 for persons with higher than secondary education, but lower than a master's degree, and 4 for those with a master's degree or higher. The average education level of employees in our sample is 2.1, which is equivalent to approximately 12.4 years of education. The diversity in education levels and fields stand at 0.564 and 0.637, respectively, indicating a higher diversity in terms of fields of education than in levels of education.

The economic activities with the highest share of firms are C (manufacturing) with 33 per cent of observations and G (wholesale and retail trade, repair of motor vehicles and motorcycles) with 23 per cent of observations. 28.4 per cent of firm-year observations are in high/medium-tech or knowledge-intensive sectors (HT/KIS).<sup>2</sup> Firms employ on average 39.9 full-time equivalent workers, and 73.5 per cent of them export their goods or services.



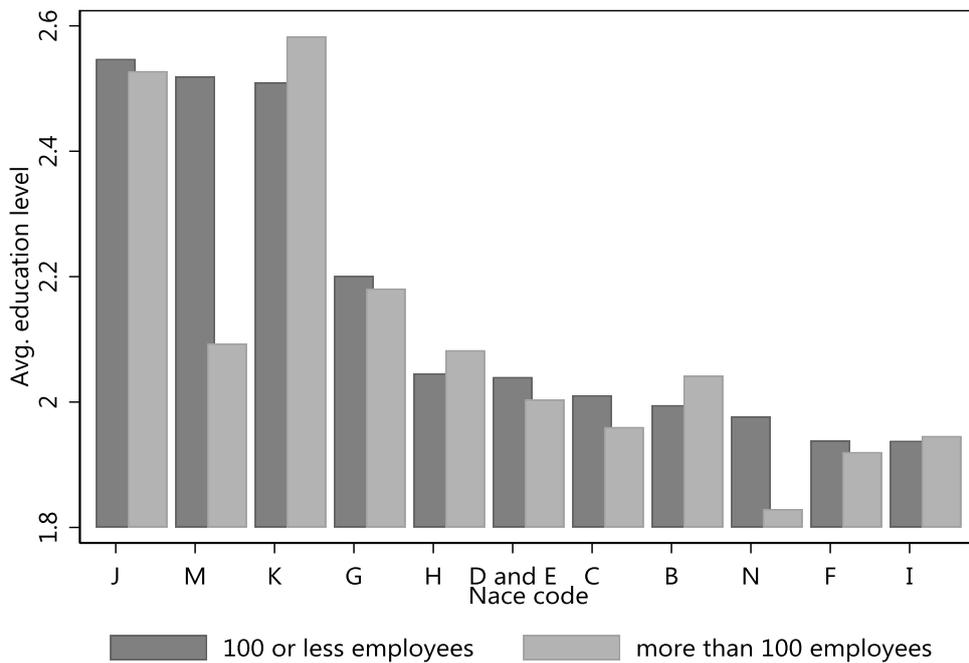
**Table 1: Descriptive statistics at the firm level (2008–2017)**

	Mean	Sd
Avg. education level	2.130	0.320
Norm. avg. education level	0.009	0.264
Diversity in education levels	0.564	0.172
Norm. diversity in education levels	0.005	0.166
Diversity in education fields	0.637	0.177
Norm. diversity in education fields	0.004	0.169
HT/KIS	0.284	0.451
Employment	39.883	40.283
ln(K/L)	9.775	1.693
ln(debt/liabilities)	-0.724	0.665
Exporter	0.735	0.442
Sales (in thousand EUR)	6,925	21,356
Foreign sales (in thousand EUR)	2,051	9,200
Nace code B	0.004	0.064
Nace code C	0.334	0.472
Nace codes D and E	0.011	0.107
Nace code F	0.104	0.306
Nace code G	0.232	0.422
Nace code H	0.064	0.244
Nace code I	0.036	0.187
Nace code J	0.059	0.236
Nace code K	0.014	0.117
Nace code M	0.098	0.297
Nace code N	0.044	0.204
Correlation[norm. diversity in education levels, norm. diversity in education fields]	0.388	
Number of observations	36,436	

Firms' human capital varies considerably among economic activities and sizes of firms, which is illustrated in Figure 1. For example, the average education level in firms with 100 employees or less within construction (Nace code F) is 1.9 (or roughly 11.6 years of education), but 2.5 (or roughly 14 years of education) in financial service activities (Nace code K). Similarly, an average employee in a firm with 100 or less employees within professional, scientific and technical activities (Nace code M) has a level of education equal to 2.5 (or roughly 14 years of education). However, an average employee's educational level in a firm with more than 100 employees within the same sector equals only 2.1 (or roughly 12.4 years of education).



**Figure 1. Average education level by Nace code and firm size**



The disparities among firms with different sizes and in various sectors are evident also in the diversity of human capital. As it is shown in Figure 2 the diversity in fields of education for firms with 100 employees or less ranges between 0.51 in financial and insurance activities (Nace code K) and 0.75 in accommodation and food service activities (Nace code I), whereas the diversity in terms of fields of education in firms with more than 100 employees in accommodation and food service activities reaches as high as 0.82. Comparable conclusions can be drawn from Figure 3, which depicts the average diversity of education levels by size and economic activity of firms.

**Figure 2. Diversity in education fields by Nace code and firm size**

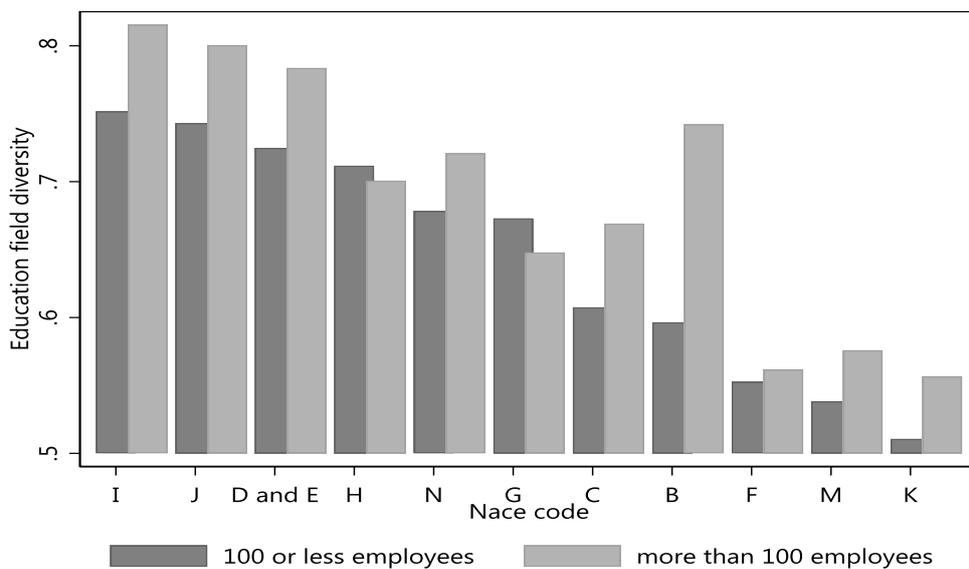
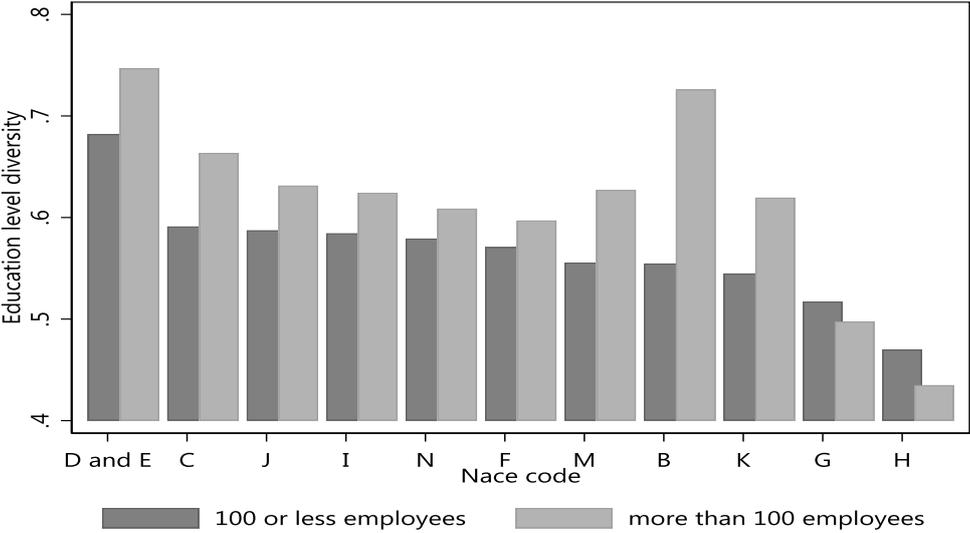


Figure 3. Diversity in education levels by Nace code and firm size



**RESULTS AND DISCUSSION**

Table 2 shows the impacts of diversity in education levels and fields (measured with normalized diversity indices) on the growth of sales and foreign sales per employee and the probability of being an exporter. Estimates indicate that an increase in average education level above competitors' (i.e. firms of similar size and in the same industry) average has a statistically significant positive effect on the growth of sales per employee in the HT/KIS sector, which is not surprising as these sectors depend on a highly educated workforce. It is also positively associated with the probability of being an exporter.

Results also reveal a different relationship between the two measures of diversity and growth of sales per employee or probability of being an exporter. While an increase of diversity in education levels above the average diversity in education levels of competitors does not have a statistically significant effect on a firm's growth of sales per employee or the likelihood of exporting, a higher diversity in education fields increases both measures of firm performance. An increase in diversity in education fields by 0.1 (or 10 percentage points) above the industry average for a particular firm size is associated with 4.1 per cent increase in growth of sales per employee and 0.35 percentage point increase in the probability of exporting for those in non-HT/KIS sector. However, the association between diversity in education fields and the likelihood of exporting is not statistically significant for firms in HT/KIS sector.



**Table 2. Estimation results**

	(1)	(2)	(3)	
	Growth of sales per employee	Growth of foreign sales per employee	Pr(exporter)	
			HT/KIS	Other sectors
Diversity in education levels	-0.130 (0.106)	-0.320 (0.435)		-0.002 (0.010)
Diversity in education fields	0.410* (0.194)	1.115 (0.632)		0.035** (0.011)
Avg. education	-0.215 (0.132)	-0.303 (0.421)		0.048*** (0.007)
HT/KIS x Diversity in education levels	0.153 (0.195)	-0.758 (0.853)	0.025 (0.017)	
HT/KIS x Diversity in education fields	-0.034 (0.294)	1.371 (1.178)	-0.017 (0.017)	
HT/KIS x Avg. education level	0.298** (0.115)	0.434 (0.420)	0.032*** (0.008)	
HT/KIS	0.097 (0.200)	0.645 (0.741)		
Arellano-Bond test for AR(2), p-value	0.475 0.075	0.063 0.067		
Sargan overidentification test p-value	4,390	3,063		
No. of firms				
Pseudo R-squared			0.652	0,659
Observations	19,983	13,797	8,733	21,772

Notes: \*  $p < 0.05$ , \*\*  $p < 0.01$ , \*\*\*  $p < 0.001$ . Clustered standard errors in parentheses. Regressions also control for: firms size, capital to labour ratio, debt to liabilities ratio, industries (10 dummies) and year dummies. Regressions for growth of sales and foreign sales per employee also contain the first lag of the dependent variable. Regressions for the probability of being exporter control for the logarithm of foreign sales in the previous period as well. Average education, diversity in education levels and fields are measured as deviation from the average of firms of similar size and the same industry. Column 3 reports average marginal effects of the probit model separately for the HT/KIS sector and other sectors due to the complexity of marginal effects of interaction terms in probit models (e.g. Ai & Norton, 2003)

The results clearly indicate that the impacts of diversity in education levels and diversity in education fields on firm performance differ in size. Their size also depends on the measure of firm performance and the firm's sector. Therefore, the lack of consensus in the existing empirical literature, which concentrates on different industries, and lacks a common measure of diversity or performance indicator, comes as a no surprise. Because the two measures are positively correlated (see Table 1), a large part of the positive effects of education diversity on firms' performance in studies measuring only diversity in terms of levels could be due to the omitted variables bias. In other words, in those studies, diversity in education levels might be actually measuring the effect of diversity in education fields.



Here presented results are in line with studies that measured the diversity in terms of field of education, and all found a positive effect on performance (Hambrick, 1996; Navon, 2010; Kearney & Gebert, 2009). They are also supported by studies that used a combined measure of education level and field diversity and estimated positive or insignificant effects on firms' performance (Parrotta et al., 2014a, 2014b; Backman & Kohlhase, 2020).

A possible explanation for the positive effects of diversity in fields of education and insignificant effects of diversity in education levels on firms' performance is Lazear's (1999) theory. It claims that firms benefit the most from labour force diversity if individuals with disjoint skills can communicate with each other. The communication among workers with different educational levels might be more prone to barriers and conflicts (Pelled & Adler, 1994; Secord & Backman, 1974), than among individuals with different specializations (Pelz, 1956), possibly due to greater differences in norms, beliefs and behaviour between than within groups of individuals with specific level but different fields of education (Gifford & Nilsson, 2014; Schommer, 1998). As a result, the labour force diversity in education fields benefits a firm's performance more than diversity in terms of levels of education.

Ilmakunnas and Ilmakunnas (2011), however, offered an alternative explanation saying that firms specialize in low or high skill tasks. According to their theory, firms with a skill-diverse workforce (in terms of levels of education) are the ones that failed to specialize and have, as a consequence, lower performance. According to here presented evidence, this explanation is less likely.

## CONCLUSIONS

The analysis of education diversity–growth nexus in firms from a wide range of economic activities emphasized the importance of differentiating between diversity in terms of levels and diversity in terms of fields of education. While employment of workers with more diverse specializations than those of competitors benefits firm's growth and extensive margin of trade (of non-HT/KIS firms), an above-average workforce heterogeneity in education levels does not seem to affect these measures of firm performance.

The paper showed that authors trying to measure the workforce's heterogeneity should bear in mind that educational level diversity and educational field diversity do not have equivalent impacts on firm performance. What is more, the two measures are positively correlated, meaning that studies concentrating on the heterogeneity of educational levels are actually estimating the combined impact of diversity in terms of levels and fields of education due to the omitted variable bias. The presented results also offer some reconciliation between previous mixed results in the literature on education diversity and firm's performance, as they show that association differs by economic activity of the firm and a measure of firms' performance.

Although the results seem to fit Lazear's (1999) theoretical model regarding the relationship between a firm's growth and labour diversity, further research with additional data on communication problems among different groups of workers is needed to support such claims fully. Unfortunately, the administrative data used in this paper does not enable such examination. Future research should also try to measure differences in the actual level and specialization of skills instead of formal education as the latter present only a proxy of an individual's actual abilities. A limitation arising from the Slovenian economy structure is the lack of large firms in our sample. The researchers should investigate whether the associations between the constructs differ in large firms compared to small and medium ones. But even though this paper left the explanation as to why diversity in education fields is and diversity in education levels is not beneficial for firms' performance to future studies, it provided support for advising firms and their managers to strive for diversity in education fields of their workforce.



## NOTES

1. The dataset includes private firms in NACE sections: B – mining and quarrying, C – manufacturing, D – electricity, gas, steam and air conditioning supply, E – water supply, sewerage, waste management and remediation activities, F – construction, G – wholesale and retail trade, repair of motor vehicles and motorcycles, H – transportation and storage, I – accommodation and food service activities, J – information and communication, K – financial and insurance activities, M – professional, scientific and technical activities, and N – administrative and support service activities.
2. According to Eurostat (2020b), high-technology firms are found in the manufacture of basic pharmaceutical products and pharmaceutical preparations (NACE 21) and manufacture of computer, electronic and optical products (NACE 26). Medium-high technology firms are those in the manufacture of chemicals and chemical products (NACE 20), manufacture of electrical equipment (NACE 27), manufacture of machinery and equipment n.e.c. (NACE 28), manufacture of motor vehicles, trailers and semi-trailers (NACE 29), and manufacture of other transport equipment (NACE 30). Knowledge-intensive services include water transport (NACE 50), air transport (NACE 51), information and communication (NACE section J), financial and insurance activities (NACE section K), professional, scientific and technical activities (NACE section M), employment activities (NACE 78), security and investigation activities (NACE 80), public administration and defence, compulsory social security (NACE section O), education (NACE section P), human health and social work activities (NACE section Q), arts, entertainment and recreation (section R).

## COMPLIANCE WITH ETHICAL STANDARDS

Conflict of Interest: The authors declare that they have no conflict of interest.

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## REFERENCES

- Ai, C. R., & E. C. Norton. (2003). Interaction terms in logit and probit models. *Economics Letters*, 80(1), 123–129.
- Arellano, M., & Bond, S. (1991). Some tests of specification for panel data: Monte Carlo evidence and an application to employment equations. *The review of economic studies*, 58(2), 277–297.
- Arellano, M., & Bover, O. (1995). Another look at the instrumental variable estimation of error-components models. *Journal of econometrics*, 68(1), 29–51.
- Backman, M., & Kohlhase, J. E. (2020). Labor force diversity and firm survival. *Journal of Regional Science*, 60(5), 903–928.
- Becker, G. S. (1964). *Human capital: A theoretical and empirical analysis, with special reference to education*. University of Chicago press.
- Bernard, A. B., & Jensen, J. B. (1999). Exceptional exporter performance: cause, effect, or both?. *Journal of international economics*, 47(1), 1–25.
- Black, S. E., & Lynch, L. M. (1996). Human-capital investments and productivity. *The American economic review*, 86(2), 263–267.
- Blundell, R., & Bond, S. (1998). Initial conditions and moment restrictions in dynamic panel data models. *Journal of econometrics*, 87(1), 115–143.
- Brambilla, I., Lederman, D., & Porto, G. (2017). Exporters, engineers, and blue-collar workers. *The World Bank Economic Review*, 30(Supplement\_1), S126–S136.
- Chiappini, R., & Jégourel, Y. (2021). "The buck stops with the executives": Assessing the impact of workforce composition and cultural distance on French firms' exports. *Economic Modelling*, 94, 45–57.
- Clerides, S. K., Lach, S., & Tybout, J. R. (1998). Is learning by exporting important? Micro-dynamic evidence from Colombia, Mexico, and Morocco. *The quarterly journal of economics*, 113(3), 903–947.
- Coad, A. (2009). *The growth of firms: A survey of theories and empirical evidence*. Edward Elgar Publishing.
- Davidsson, P., Achtenhagen, L., & Naldi, L. (2010). *Small firm growth*. Now Publishers Inc.
- Eliasson, K., Hansson, P., & Lindvert, M. (2012). Do firms learn by exporting or learn to export? Evidence from small and medium-sized enterprises. *Small Business Economics*, 39(2), 453–472.
- Eurostat. (2020a). *Eurostat database*. <https://ec.europa.eu/eurostat/data/database>. (Accessed: 1. 9. 2020)
- Eurostat. (2020b). *Eurostat indicators on high-tech industry and knowledge-intensive services*. [https://ec.europa.eu/eurostat/cache/metadata/Annexes/htec\\_esms\\_an3.pdf](https://ec.europa.eu/eurostat/cache/metadata/Annexes/htec_esms_an3.pdf). (Accessed: 2. 9. 2020)
- Festinger, L. (1954). A theory of social comparison processes. *Human relations*, 7(2), 117–140.
- Ganotakis, P., & Love, J. H. (2011). R&D, product innovation, and exporting: evidence from UK new technology based firms. *Oxford Economic Papers*, 63(2), 279–306.
- Garnero, A., Kampelmann, S., & Rycx, F. (2014). The heterogeneous effects of workforce diversity on productivity, wages, and profits. *Industrial Relations: A Journal of Economy and Society*, 53(3), 430–477.
- Gifford, R., & Nilsson, A. (2014). Personal and social factors that influence proenvironmental concern and behaviour: A review. *International Journal of Psychology*, 49(3), 141–157.
- Grund, C., & Westergaard-Nielsen, N. (2008). Age structure of the workforce and firm performance. *International Journal of Manpower*.



- Hamilton, B. H., Nickerson, J. A., & Owan, H. (2012). Diversity and productivity in production teams. In *Advances in the economic analysis of participatory and labor-managed firms*. Emerald Group Publishing Limited.
- Hambrick, D. C., Cho, T. S., & Chen, M. J. (1996). The influence of top management team heterogeneity on firms' competitive moves. *Administrative science quarterly*, 659-684.
- Hirsch, B., Jahn, E.J. and Zwick, T. (2020), Birds, Birds, Birds: Co-Worker Similarity, Workplace Diversity and Job Switches. *British Journal of Industrial Relations*, 58, 690-718.
- Ilmakunnas, P., & Ilmakunnas, S. (2011). Diversity at the workplace: whom does it benefit? *De Economist*, 159(2), 223-255.
- Iranzo, S., Schivardi, F., & Tosetti, E. (2008). Skill dispersion and firm productivity: An analysis with employer-employee matched data. *Journal of Labor Economics*, 26(2), 247-285.
- Jackson, S. E., Brett, J. F., Sessa, V. I., Cooper, D. M., Julin, J. A., & Peyronnin, K. (1991). Some differences make a difference: Individual dissimilarity and group heterogeneity as correlates of recruitment, promotions, and turnover. *Journal of applied psychology*, 76(5), 675.
- Jehn, K. A. (1995). A multimethod examination of the benefits and detriments of intragroup conflict. *Administrative science quarterly*, 256-282.
- Jehn, K. A., & Bezrukova, K. (2004). A field study of group diversity, workgroup context, and performance. *Journal of Organizational Behavior*, 25(6), 703-729.
- Jehn, K. A., Northcraft, G. B., & Neale, M. A. (1999). Why differences make a difference: A field study of diversity, conflict and performance in workgroups. *Administrative science quarterly*, 44(4), 741-763.
- Kearney, E., & Gebert, D. (2009). Managing diversity and enhancing team outcomes: the promise of transformational leadership. *Journal of applied psychology*, 94(1), 77.
- Knight, D., Pearce, C. L., Smith, K. G., Olian, J. D., Sims, H. P., Smith, K. A., & Flood, P. (1999). Top management team diversity, group process, and strategic consensus. *Strategic management journal*, 20(5), 445-465.
- Kurtulus, F. A. (2011). What types of diversity benefit workers? empirical evidence on the effects of co-worker dissimilarity on the performance of employees. *Industrial Relations: A Journal of Economy and Society*, 50(4), 678-712.
- Lazear, E. P. (1999). Globalization and the market for team-mates. *The Economic Journal*, 109(454), 15-40.
- Lee, H. W., & Kim, E. (2020). Workforce diversity and firm performance: Relational coordination as a mediator and structural empowerment and multisource feedback as moderators. *Human Resource Management*, 59(1), 5-23.
- Luong, T.A. and Chen, W.-C. (2016), Margins of Exports and Labor. *Review of Development Economics*, 20, 201-213.
- Melitz, M. J. (2003). The impact of trade on intra-industry reallocations and aggregate industry productivity. *Econometrica*, 71(6), 1695-1725.
- Mulliqi, A., Adnett, N., & Hisarciklilar, M. (2019). Human capital and exports: A micro-level analysis of transition countries. *The Journal of International Trade & Economic Development*, 28(7), 775-800.
- Munch, J. R., & Skaksen, J. R. (2008). Human capital and wages in exporting firms. *Journal of International Economics*, 75(2), 363-372.
- Navon, G. (2010). Human capital spillovers in the workplace: Labor diversity and productivity. *Israel Economic Review*, 8(1), 69-90.



- Parrotta, P., Pozzoli, D., & Pytlikova, M. (2014a). Labor diversity and firm productivity. *European Economic Review*, 66, 144–179.
- Parrotta, P., Pozzoli, D., & Pytlikova, M. (2014b). The nexus between labor diversity and firm's innovation. *Journal of Population Economics*, 27(2), 303–364.
- Pelled, L. H., & Adler, P. S. (1994). Antecedents of intergroup conflict in multifunctional product development teams: A conceptual model. *IEEE transactions on engineering management*, 41(1), 21–28.
- Pelz, D. C. (1956). Some social factors related to performance in a research organization. *Administrative Science Quarterly*, 310–325.
- Pfeffer, J. (1985). Organizational demography: Implications for management. *California management review*, 28(1), 67–81.
- Prat, A. (2002). Should a team be homogeneous? *European Economic Review*, 46(7), 1187–1207.
- Schommer, M. (1998). The influence of age and education on epistemological beliefs. *British Journal of Educational Psychology*, 68(4), 551–562.
- Secord, P. F., & Backman, C. W. (1974). *Social psychology 2nd edition*. New York: McGraw-Hill.
- Simons, T., Pelled, L. H., & Smith, K. A. (1999). Making use of difference: Diversity, debate, and decision comprehensiveness in top management teams. *Academy of management journal*, 42(6), 662–673.
- Smith, K. G., Smith, K. A., Olian, J. D., Sims Jr, H. P., O'Bannon, D. P., & Scully, J. A. (1994). Top management team demography and process: The role of social integration and communication. *Administrative science quarterly*, 412–438.
- Stewart, G. L. (2006). A meta-analytic review of relationships between team design features and team performance. *Journal of management*, 32(1), 29–55.
- SURS. (2020). *Sistat database*. <https://pxweb.stat.si/SiStat/en>. (Accessed: 1. 9. 2020)
- Tsui, A. S., Egan, T. D., & O'Reilly III, C. A. (1992). Being different: Relational demography and organizational attachment. *Administrative science quarterly*, 549–579.
- Vandenberghe, V. (2016). Is workforce diversity good for efficiency? an approach based on the degree of concavity of the technology. *International Journal of Manpower*, 37(2), 253–267.
- Wagner, J. (2012). Average wage, qualification of the workforce and export performance in German enterprises: evidence from KombiFiD data. *Journal for Labour Market Research*, 45(2), 161–170.

